

Méthode d'ensemble pour l'estimation des variances d'erreur d'ébauche dans un système 3D-Var appliqué à l'océan global

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- 1 General context and objectives
- 2 The assimilation system
- 3 Background-error variance specification
- 4 Numerical experiments
- 5 Conclusions and future work

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- This study employs a variational data assimilation system (OPAVAR) developed at CERFACS for climate research applications.
- The system has been applied in the European projects ENACT and ENSEMBLES:
 - ▶ to provide initial conditions for climate hindcasts/forecasts with coupled ocean-atmosphere models;
 - ▶ to reconstruct the 1960-present history of the ocean state (“reanalysis”).
- In ENSEMBLES, a 3D-Var version of OPAVAR was used to produce a 9-member ensemble of multi-decadal ocean analyses:
 - ▶ produced using perturbed atmospheric surface fluxes;
 - ▶ used to sample uncertainty in ocean initial conditions for probabilistic forecasts on seasonal to decadal time-scales.

- An important feature of an ensemble data assimilation system is its capacity to provide flow-dependent information on analysis and background error.
- This information can be exploited to improve the estimate of the background-error covariance matrix (**B**) on each assimilation cycle.
 - ▶ In the ENSEMBLES experiments, we made no attempt to use the ensemble to update **B**.
- The objective here is explore the possibility of using the ENSEMBLES data assimilation system to improve **B**.

- Construct a low-rank approximation to \mathbf{B} directly from the sample covariance of the ensemble of model forecast states.
(Houtekammer and Mitchell 2001; Keppenne and Reinecker 2002; Ott *et al.* 2004; Buehner and Charron 2007; Oke *et al.* 2007).
 - ▶ Covariance localization is necessary to minimize spurious effects due to sampling error.

- or -

- Use the ensemble indirectly to calibrate parameters of a (localized) covariance model in a full-rank (operator) representation of \mathbf{B} .
(Fisher 2003; Žagar *et al.* 2005; Belo Pereira and Berre 2006; Berre *et al.* 2006; Küçükkaraca and Fisher 2006).
 - ▶ A flexible covariance model (inhomogeneous, anisotropic) is required to make best use of the ensemble information.
 - ▶ Full-rank covariance matrix representations are important for fitting detailed structures in the observations.

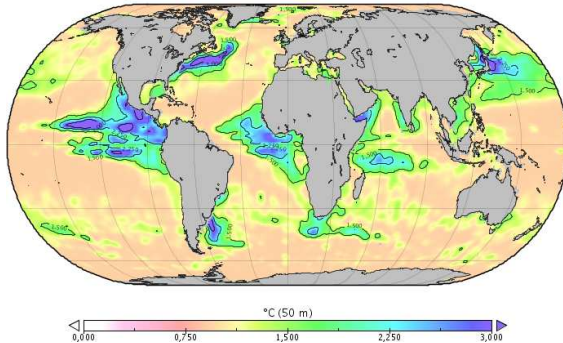
- In this study, we adopt the covariance model approach.
- In particular, we investigate the potential of an ensemble of ocean states to provide useful flow-dependent estimates of the background-error **variances** in the OPAVAR 3D-Var system.
- This approach will be compared with a simpler approach for incorporating flow dependence in the variances.
- This study is a first step towards making more comprehensive use of an ensemble for calibrating additional parameters of the covariance model.

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- The ocean model is the global ORCA2 configuration of OPA8.2 (Madec *et al.* 1998).
- The surface forcing fields are derived from ERA40 (Uppala *et al.* 2005).
- The assimilation method is a multivariate 3D-Var version of OPAVAR (Weaver *et al.* 2005).
- First-Guess at Appropriate Time (FGAT) and Incremental Analysis Updates (IAU) are employed.
- The data are quality-controlled temperature and salinity profiles from EN2v1 data-base (Ingleby and Huddleston 2007).
- The observation-error variances are estimated from a model-data comparison prior to assimilation (the Fu *et al.* (1993) method).

Example of temperature σ^o at 50 m

Ecarts-types d'erreur d'observation de température



$$J[\delta\mathbf{w}] = \frac{1}{2}\delta\mathbf{w}^T \mathbf{B}_{(\mathbf{w})}^{-1} \delta\mathbf{w} + \frac{1}{2}(\mathbf{H}\delta\mathbf{w} - \mathbf{d})^T \mathbf{R}^{-1}(\mathbf{H}\delta\mathbf{w} - \mathbf{d})$$

where

$$\mathbf{d} = \begin{pmatrix} \mathbf{d}_0 \\ \vdots \\ \mathbf{d}_i \\ \vdots \\ \mathbf{d}_N \end{pmatrix} = \begin{pmatrix} \mathbf{y}_0^o - \mathbf{H}_0 \mathbf{w}^b(t_0) \\ \vdots \\ \mathbf{y}_i^o - \mathbf{H}_i \mathbf{w}^b(t_i) \\ \vdots \\ \mathbf{y}_N^o - \mathbf{H}_N \mathbf{w}^b(t_N) \end{pmatrix} \quad \text{and} \quad \mathbf{H} = \begin{pmatrix} \mathbf{H}_0 \\ \vdots \\ \mathbf{H}_i \\ \vdots \\ \mathbf{H}_N \end{pmatrix}.$$

- $\delta\mathbf{w} = (\delta T, \delta S)^T$ is the vector of temperature and salinity increments.
- $\mathbf{y}_i^o = (T_i^o, S_i^o)^T$ is the vector of temperature and salinity observations.
- Increments for sea-surface height and velocity are obtained *a posteriori* using balance constraints applied to the analysis increment $\delta\mathbf{w}^a$.

$$\mathbf{B}_{(\mathbf{w})} = \mathbf{K}_{(\mathbf{w})} \mathbf{D}_{(\hat{\mathbf{w}})}^{1/2} \mathbf{F}_{(\hat{\mathbf{w}})} \mathbf{F}_{(\hat{\mathbf{w}})}^T \mathbf{D}_{(\hat{\mathbf{w}})}^{1/2} \mathbf{K}_{(\mathbf{w})}^T$$

where

$$\mathbf{F}_{(\hat{\mathbf{w}})} = \begin{pmatrix} \mathbf{F}_{TT} & 0 \\ 0 & \mathbf{F}_{S_U S_U} \end{pmatrix}, \quad \mathbf{D}_{(\hat{\mathbf{w}})}^{1/2} = \begin{pmatrix} \mathbf{D}_T^{1/2} & 0 \\ 0 & \mathbf{D}_{S_U}^{1/2} \end{pmatrix}, \quad \mathbf{K}_{(\mathbf{w})} = \begin{pmatrix} \mathbf{I} & 0 \\ \mathbf{K}_{ST} & \mathbf{I} \end{pmatrix}$$

- $\hat{\mathbf{w}} = (T, S_U)^T$ where S_U corresponds to “unbalanced” salinity.
- $\mathbf{K}_{(\mathbf{w})}$ is a multivariate balance operator: $\hat{\mathbf{w}} \mapsto \mathbf{w}$.
- $\mathbf{F}_{(\hat{\mathbf{w}})} \mathbf{F}_{(\hat{\mathbf{w}})}^T$ is a quasi-Gaussian 3D univariate correlation operator, modelled using a diffusion operator (see Mirouze and Weaver poster).
- $\mathbf{D}_{(\hat{\mathbf{w}})}$ is a variance matrix (for $\hat{\mathbf{w}}$), whose estimation is the focus of this study.

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For temperature:

$$\sigma_T^b = \begin{cases} \max(\tilde{\sigma}_T^b, \sigma_T^{ml}) & \text{in the mixed layer,} \\ \max(\tilde{\sigma}_T^b, \sigma_T^{do}) & \text{below the mixed layer,} \end{cases}$$

where

$$\tilde{\sigma}_T^b = \min(|(\partial T / \partial z|_{T=T^b}) \delta z|, \sigma_T^{max}).$$

For unbalanced salinity:

$$\sigma_{S_U}^b = \begin{cases} \sigma_{S_U}^{max} & \text{if } z \geq z_{max} \\ \sigma_{S_U}^{max} \alpha(z) & \text{if } z < z_{max} \end{cases}$$

where z_{max} is the depth of maximum $|(\partial S / \partial T|_{T=T^b})|$ and

$$\alpha(z) = 0.1 + 0.45 \times \{1 - \tanh(2 \ln(z/z_{max}))\}.$$

Advantages:

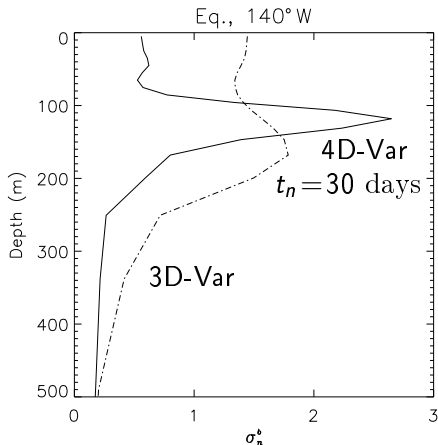
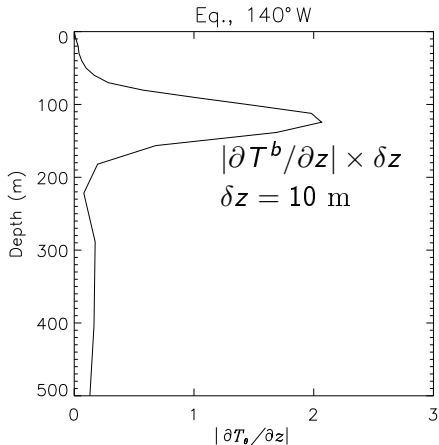
- A simple and inexpensive way to incorporate flow dependence in σ^b .
- Captures some of the dynamical effects implicit in 4D-Var (Weaver *et al.* 2003).

Disadvantages:

- Somewhat adhoc (especially for $\sigma_{S_U}^b$), and requires careful tuning of several parameters (δz , σ_T^{ml} , σ_T^{do} , σ_T^{max} , $\sigma_{S_U}^{max}$).
- Does not account for changes in the accuracy of σ^b resulting from data assimilation on previous assimilation cycles (i.e., σ^b is independent of the observing system).

Implicitly-evolved temperature σ^b in 4D-Var

(from Weaver *et al.* 2003)



$$\mathbf{P}^b(t_n) = \mathbf{B} \quad \text{in 3D-Var FGAT}$$

$$\mathbf{P}^b(t_n) = \mathbf{M}(t_0, t_n) \mathbf{B} \mathbf{M}(t_0, t_n)^T \quad \text{in 4D-Var (cf. EKF)}$$

Estimate σ^b from the difference between background states of successive ensemble members, $l = 0, \dots, L - 1$:

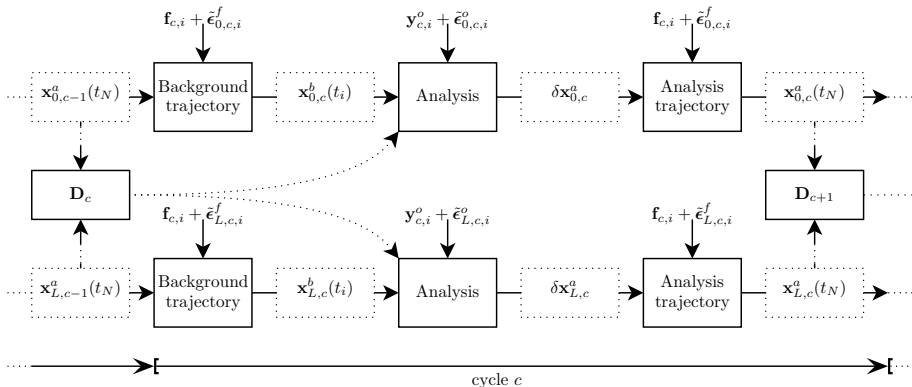
$$\begin{aligned} \mathbf{D}_{(\hat{\mathbf{w}})} &= \text{diag} \left\{ (\sigma_T^b)^2, (\sigma_{S_U}^b)^2 \right\} \\ &= \text{diag} \left\{ \frac{1}{2(L-1)} \sum_{l=0}^{L-1} \left[\mathbf{K}_{(\mathbf{w})}^{-1} \left(\mathbf{w}_l^b(t_0) - \mathbf{w}_{l+1}^b(t_0) \right) \right] \right. \\ &\quad \left. \times \left[\mathbf{K}_{(\mathbf{w})}^{-1} \left(\mathbf{w}_l^b(t_0) - \mathbf{w}_{l+1}^b(t_0) \right) \right]^T \right\} \end{aligned}$$

where

$$\mathbf{K}_{(\mathbf{w})}^{-1} = \begin{pmatrix} \mathbf{I} & 0 \\ -\mathbf{K}_{ST} & \mathbf{I} \end{pmatrix}$$

and $\mathbf{w}_L^b(t_0) = \mathbf{w}_0^b(t_0)$.

The ensemble 3D-Var cycling procedure



- The background-error variance matrix (\mathbf{D}_c) used for the analysis on cycle c is estimated from the sample variance matrix computed from the ensemble of background states ($\mathbf{x}_{l,c}^b(t_0)$) at the start of cycle c .
- In our set-up, $\mathbf{x}_{l,c}^b(t_0) = \mathbf{x}_{l,c-1}^a(t_N)$.

Advantages:

- The analysis and forecast state perturbations from the cycled ensemble system will evolve like the true analysis and forecast errors. (Berre *et al.* 2006; Daget *et al.* 2008)
- If the perturbations to the system input parameters (observations, initial state, model error, forcing and boundary terms) are random samples drawn from the pdf of the true errors then

$$\begin{aligned}
 \mathbf{P}^b(t_0) &\equiv E[(\mathbf{x}^b(t_0) - \mathbf{x}^t(t_0))(\mathbf{x}^b(t_0) - \mathbf{x}^t(t_0))^T] \\
 &\approx \frac{1}{L-1} \sum_{l=1}^{L-1} (\mathbf{x}_l^b(t_0) - \mathbf{x}^b(t_0)) (\mathbf{x}_l^b(t_0) - \mathbf{x}^b(t_0))^T \\
 &\approx \frac{1}{2(L-1)} \sum_{l=0}^{L-1} (\mathbf{x}_l^b(t_0) - \mathbf{x}_{l+1}^b(t_0)) (\mathbf{x}_l^b(t_0) - \mathbf{x}_{l+1}^b(t_0))^T
 \end{aligned}$$

where $\mathbf{x}_L^a(t_0) = \mathbf{x}_0^a(t_0) = \mathbf{x}^a(t_0)$.

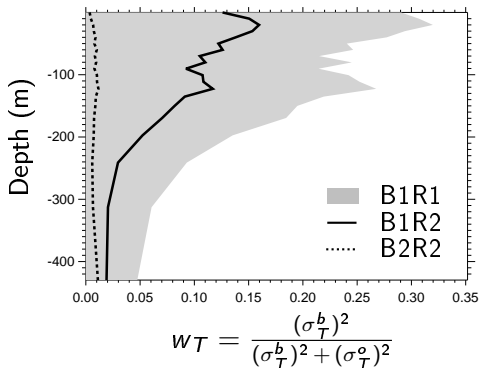
The disadvantages are essentially practical.

- An ensemble-based data assimilation system is expensive.
 - ▶ A model integration and variational analysis must be performed separately for each ensemble member.
- Designing an effective perturbation strategy is tricky.
 - ▶ Which input parameters to perturb?
 - ▶ How to obtain an adequate spread with few ensemble members?
- Constructing appropriate perturbations to the system input parameters is difficult.
 - ▶ We can only approximately sample the pdf of the true errors.

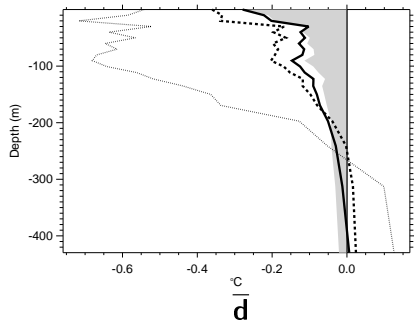
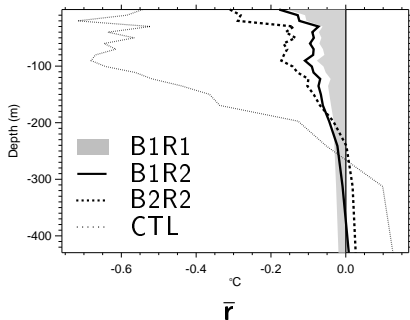
- A 9-member ensemble.
- The perturbed input parameters:
 - ▶ the surface forcing fields (heat flux, fresh-water flux, wind-stress);
 - ▶ the temperature and salinity observations;
 - ▶ the background state;
 - ▶ model error is neglected.
- Construction of the perturbations:
 - ▶ the forcing perturbations are derived from differences between different forcing analysis products (Balmaseda *et al.* 2008);
 - ▶ the observation perturbations are drawn from a Gaussian pdf with covariance matrix \mathbf{R} ;
 - ▶ the background state is perturbed implicitly via the cycling procedure;
- Reduction of sampling error:
 - ▶ A 90-day (9-cycle) sliding window is used, giving an effective ensemble size of 81 on each cycle for estimating σ^b .
 - ▶ Intraseasonal variability in σ^b is thus filtered out.

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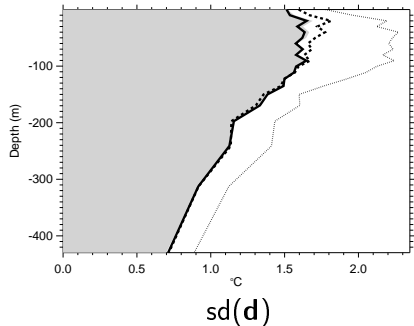
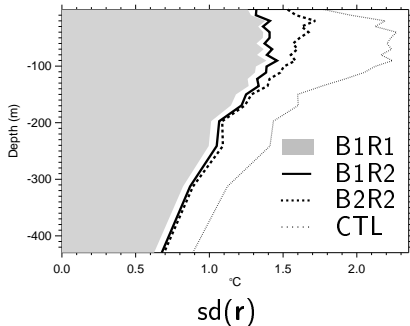
- The experimental design follows the common reanalysis procedures used in the ENSEMBLES and ENACT projects (Davey *et al.* 2006).
- The experiments are performed for the 9-year period from 1 January 1993 to 31 December 2001.
- A 10-day assimilation cycle is used.
- The experiments:
 - ▶ **CTL** : no data assimilation.
 - ▶ **B1R1** : parameterized σ^b , and σ^o defined using globally-averaged estimates from Ingleby and Huddleston (2007).
 - ▶ **B1R2** : parameterized σ^b , and σ^o estimated from Fu *et al.* method.
 - ▶ **B2R2** : ensemble σ^b , and σ^o estimated from Fu *et al.* method.
- Results will be displayed for temperature only (results for salinity are qualitatively similar).



- Neglecting correlations, w_T is the average weight for an innovation.
- Both σ_T^b and σ_T^o have been computed at observation points, and averaged over the 1994-2000 period and the global domain.

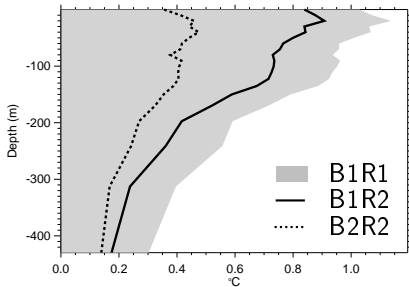


- $\mathbf{r} = \mathbf{y}^o - \mathbf{Hw}^a$ (residual) and $\mathbf{d} = \mathbf{y}^o - \mathbf{Hw}^b$ (innovation).
- \bar{z} indicates spatial (global) and temporal (1994-2000) average.
- Mean bias in CTL is reduced substantially in all assimilation expts.

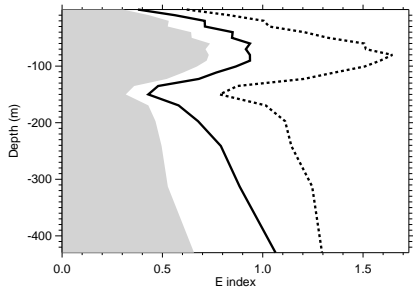


- $\mathbf{r} = \mathbf{y}^o - \mathbf{H}\mathbf{w}^a$ (residual) and $\mathbf{d} = \mathbf{y}^o - \mathbf{H}\mathbf{w}^b$ (innovation).
- $sd(\mathbf{z}) = \sqrt{(\mathbf{z} - \bar{\mathbf{z}})^2}$
- All assimilation expts. improve the fit to the observed variability.
- The “error growth” in the 10-day forecast is smallest for B2R2.

An “efficiency” index for the assimilation method

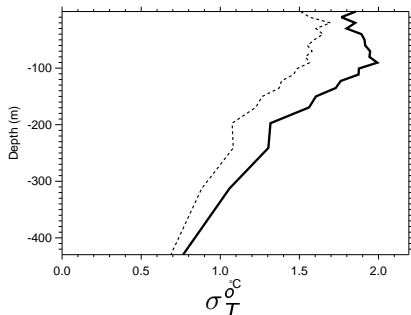
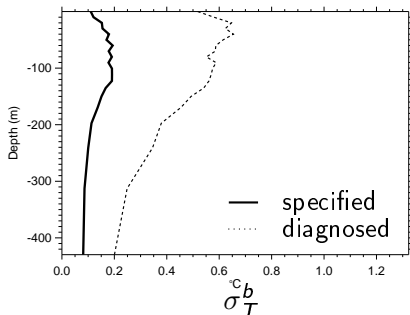


$\text{rms}(\mathbf{H}\delta\mathbf{w}^a)$



$$E = \frac{\text{rms}(\mathbf{d}_{\text{CTL}}) - \text{rms}(\mathbf{d})}{\text{rms}(\mathbf{H}\delta\mathbf{w}^a)}$$

- $E = \frac{\text{10-day forecast error improvement with respect to CTL}}{\text{“work done” by assimilation method to reduce forecast error}}$
- $E > 0$ ($E < 0$) implies assimilation is beneficial (detrimental).
- E increases (decreases) if \mathbf{d} or $\delta\mathbf{w}^a$ decreases (increases).

(method of Desroziers *et al.* 2005)

- If **B** and **R** are good estimates of the true background- and observation-error covariance matrices then

$$E[d(\mathbf{H}\delta\mathbf{w}^a)^T] \approx \mathbf{H}\mathbf{B}_{(\mathbf{w})}\mathbf{H}^T$$

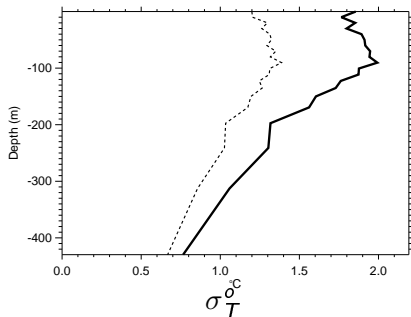
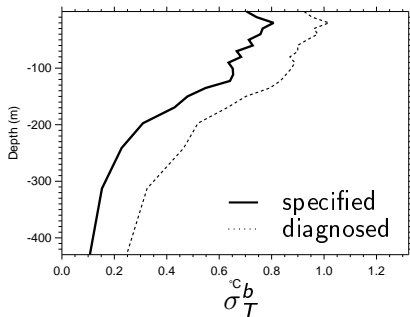
$$E[d(d - \mathbf{H}\delta\mathbf{w}^a)^T] \approx \mathbf{R}$$

- σ_T^b is **underestimated**, and σ_T^o is **overestimated**.

Specified versus diagnosed σ^b and σ^o for temperature in B1R2

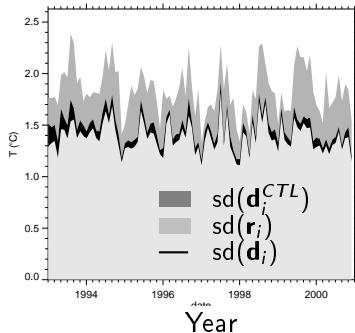
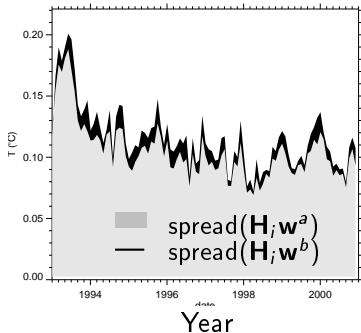


(method of Desroziers *et al.* 2005)



- σ_T^b is **underestimated** (to a lesser extent than in B2R2).
- σ_T^o is **overestimated** (to a greater extent than in B2R2).

Experiment B2R2

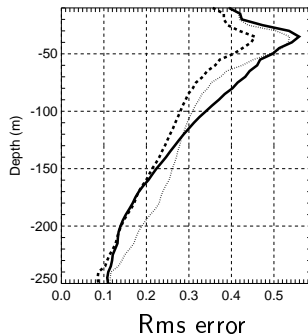
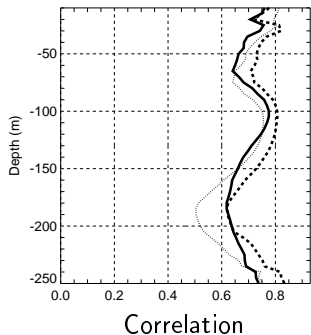


- $\text{spread}\{\mathbf{H}; \mathbf{w}^{a,b}\} = \sqrt{\frac{1}{L-1} \sum_{l=0}^{L-1} \left(\mathbf{H}_i \mathbf{w}_l^{a,b}(t_i) - \frac{1}{L} \sum_{l=0}^{L-1} \mathbf{H}_i \mathbf{w}_l^{a,b}(t_i) \right)^2}$
- Spread of the analysis < spread of the background.
- No evidence of ensemble collapse.
- $\text{Spread}(\mathbf{H}; \mathbf{w}^{a,b})$ is approximately a factor 10 smaller than $sd(\mathbf{r}_i)$, $sd(\mathbf{d}_i)$.

Comparisons with independent data

Current-meter data from TAO moorings

(Example from the eastern Pacific (110°W))



— B1R2
 B2R2
 - - - CTL

- B2R2 outperforms B1R2 (and B1R1) at all moorings.
- B2R2 outperforms CTL in the central and eastern Pacific, but slightly worse in the western Pacific.

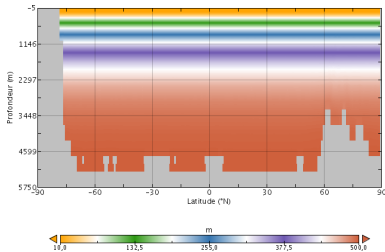
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- Both the parameterized and ensemble σ^b formulations produce a significant reduction in the rms of the innovations, with the parameterized σ^b slightly better above 150 m.
- The ensemble σ^b analyses are better “balanced” and closer to independent data (sea-level anomalies from T/P and current-meter data from TAO) than those produced with the parameterized σ^b .
- Statistical consistency checks suggest that the ensemble σ^b are underestimated. The parameterized σ^b are also underestimated but to a lesser extent.
- The apparent underestimation of the ensemble spread points to the need to increase the ensemble size and/or to improve the ensemble generation strategy.
- Producing an ensemble of analyses is costly but may be justified if these analyses can be used simultaneously for probabilistic forecasting as well as background-error covariance estimation.

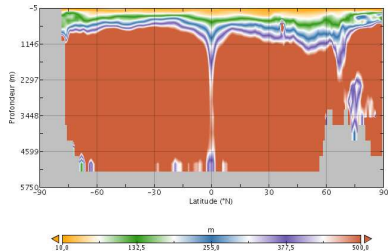
- Implementation of the ensemble ocean analysis procedure in the NEMOVAR system (OPAVAR's successor) and application to higher resolution global configurations.
 - ▶ A contribution to the NEMOVAR project, supported by LEFE-ASSIM.
- Improvements and extensions to the ensemble generation strategy (e.g., to include explicit model-error perturbations).
- Exploitation of the ensemble information for estimating other aspects of the background-error covariance matrix
 - ▶ Geographically-dependent length scales for our diffusion-based correlation model.
(Pannekoucke *et al.* 2008; Daget 2008, *PhD thesis*)
 - ▶ Testing the validity of assumptions made in the balance operator.
(Daget 2008)

Postdoctoral position available at CERFACS: see
http://www.cerfacs.fr/Open_Positions/index_offres.php

Example for temperature (from Daget 2008, PhD thesis)



Fixed scales currently used



Ensemble estimates